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# Evaluating the recruitment process into UK anaesthesia core training: a national data linkage study of doctors' performance at selection and subsequent postgraduate training

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## Postgraduate Medical Journal

### Evaluating the recruitment process into UK anaesthesia core training: doctors' performance at selection and subsequent postgraduate training- a national data linkage study

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3 **Evaluating the recruitment process into UK anaesthesia core training: doctors'**  
4 **performance at selection and subsequent postgraduate training- a national data**  
5 **linkage study**  
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## Abstract

**Purpose of the study:** To explore which factors increase the likelihood of being deemed appointable to core anaesthesia training in the UK and whether those factors subsequently predict performance in postgraduate training.

**Study design:** Observational study linking UK medical specialty recruitment data with postgraduate educational performance, as measured by Annual Review of Competence Progression (ARCP) outcomes. Data were available for 2782 trainee doctors recruited to anaesthesia core training from 2012 to 2016 with at least one subsequent ARCP outcome.

**Results:** Both higher interview and shortlisting scores were independent and statistically significant ( $p \leq 0.001$ ) predictors of more satisfactory ARCP outcomes, even after controlling for the influence of postgraduate exam failure. It was noted that a number of background variables (e.g. age at application) were independently associated with the odds of being deemed appointable at recruitment. Of these, increasing age and experience were also negative predictors of subsequent ARCP rating. These influences became statistically non-significant once ARCP outcomes associated with exam failure were excluded.

**Conclusions:** The predictors of 'appointability' largely also predict subsequent performance in postgraduate training, as indicated by ARCP ratings. This provides evidence for the validity of the selection process. Our results also suggest that greater weight could be applied to shortlisting scores within the overall process of ranking applicants for posts.

### What is already known on the subject

- Validity evidence in relation to recruitment to other medical specialties suggests that ratings on selection assessments are largely predictive of future performance
- One regionally-based UK study of selection into anaesthetics training reported similar findings, providing some evidence for the validity of the recruitment process, though national studies are lacking

### Main messages

- In recruitment to anaesthetics core training shortlisting and interview scores were strongly predictive of postgraduate performance, as measured by Annual Review of Competence Progression (ARCP) outcomes.
- Our findings support the effectiveness of the UK selection process for anaesthesia core training.

## Introduction

How medical trainees are selected into training programmes influences the quality of future consultants. Ideally, applicants should be selected on those factors that subsequently predict clinical competency. In the UK, the anaesthetics recruitment process (summarised in Figure 1) is coordinated by the Anaesthetic National Recruitment Office. Successful applicants then embark on their anaesthetics training (Figure 2).

INSERT FIGURE 1 HERE

INSERT FIGURE 2 HERE

Several studies have investigated how selection scores correlate with subsequent performance<sup>1-6</sup>. Several studies relating to General Practice have shown that performance during the selection process predicts later educational attainment<sup>7-11</sup>. One UK, regionally-based, study of selection into anaesthetics training reported a similar picture<sup>12</sup>. However, national studies are currently lacking.

There are a small number of national assessments in UK postgraduate training, including the Annual Review of Competence Progression (ARCP). This process is used to decide how a doctor progresses through postgraduate training and involves a panel of assessors reviewing a portfolio of evidence<sup>13</sup>. This includes supervisor feedback and workplace based assessments. Success also requires trainees to pass the Primary Fellowship of the Royal College of Anaesthetists exams (Primary FRCA). These exams consist of two, separately taken, sections. The first section consists of a multiple choice based knowledge test, and the second is a clinically orientated test involving an Objective Structured Clinical Examination (OSCE) and a Structured Oral Exam (SOE)<sup>14</sup>. Both of these sections must be passed during core training in order to progress to higher specialist training. Failure in these exams can

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2  
3 lead to an unsatisfactory ARCP outcome, where extended training time may be required, or  
4  
5 the trainee may be released from programme.  
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8  
9 If the selection measures are reasonably valid then performance on such metrics should also  
10  
11 predict subsequent outcomes in training, taking into account the impact of selecting out  
12  
13 unsuccessful candidates. The role of influences on both performance at recruitment and in  
14  
15 training should be taken into account where possible. These factors could include age <sup>15</sup>,  
16  
17 gender <sup>16-19</sup>, country of primary qualification <sup>20</sup>, and ethnicity <sup>21-24</sup>.  
18  
19

20  
21 This study aimed to assess the validity of the anaesthetics core training recruitment process  
22  
23 by:

- 24  
25 • Investigating the predictors of a candidate being deemed appointable to  
26  
27 anaesthetics core training  
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- 30  
31 • Evaluating the predictors of ARCP outcomes for those who enter postgraduate  
32  
33 anaesthetics training  
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- 36  
37 • Comparing any patterns observed above in order to assess the effectiveness of  
38  
39 selection into postgraduate training.  
40

## 41 **Methods**

### 42 *Data sources and preparation*

43  
44 We obtained selection data for applicants to anaesthetics core training during 2012 to 2016  
45  
46 from the Oriel database supplied to the General Medical Council (GMC). Flow of data  
47  
48 through the study is depicted in Figure 3. Our outcome measure for recruitment was  
49  
50 'deemed appointable' rather than 'appointed' in order to reduce the effect of differing  
51  
52 competition ratios across different deaneries. Shortlisting and interview performance scores  
53  
54 were standardised as z-scores by year and selection centre (i.e. transformed to have a mean  
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of zero and standard deviation of one). This controlled for differences across selection centres and for time. Thus, if doctors had applied more than once in different recruitment years, the mean interview and shortlisting scores for that particular cohort were used to standardise the scores.

INSERT FIGURE 3 HERE

Data for ARCP outcomes for trainees across the UK were available from Health Education England. This was linked to a database supplied by the GMC that included socio-demographic information, and recruitment data. To maintain anonymity the linkage was performed, via the doctors' unique registration number, by the GMC. Only 'competency-based' ARCP outcomes were included (for example, those indicating 'out of programme experience' were excluded). Anaesthetic trainees were identified by their specialty, and those in ACCS Anaesthetics and CAT were coded separately to allow for comparison between the schemes.

ARCP ratings were collapsed and recoded to form a four point ordinal (ordered categorical) scale as follows: 4 = satisfactory, 3 = additional evidence required, 2 = targeted training required but no extra time, 1 = extended training time required/left programme. This allowed the ARCP outcomes to be treated as ordinal, rather than binary (satisfactory/unsatisfactory) or nominal (unordered) in nature. This means that the information contained within the ordering of the (recoded) ARCP outcomes is preserved, increasing study power. Note that, it was important to combine the categories 'extended training time' and 'left programme'. This is because previously it has been shown that, within the context of multilevel modelling (i.e. multiple ARCPs nested within doctors), these two categories cannot be distinguished. This is assumed to be because the need to leave a programme is almost always preceded by a period of extended training time. In addition, unlike Pyne and Ben-Shlomo<sup>15</sup> we included 'insufficient evidence provided' as an



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2  
3 intermediate category of outcome. This is because it was previously noted that this outcome  
4 was associated with other undesirable outcomes (compared to a 'satisfactory' rating) and  
5 therefore was assumed to contain, on average, some information on a trainee. Thus, this  
6 approach to recoding and modelling ARCP outcomes has previously been found to be valid,  
7 and also leads to models that, at least approximately, fulfil the 'parallel odds' assumption  
8 that underlies ordinal logistic regression <sup>25</sup>.  
9

10 We included gender, age (on application and at ARCP), years of NHS experience and self-  
11 reported ethnicity (dichotomised as 'White' or 'Black and Minority Ethnic' (BME)) in  
12 analyses. We also hypothesised that being a national of, or having trained in, a country with  
13 fewer resources might also affect performance at selection or in training. Therefore we  
14 controlled for the Gross Domestic Product per capita in US Dollars (GDP) of the relevant  
15 countries of nationality and qualification, using 2008 World Bank data <sup>26</sup>.  
16  
17

### 18 *Statistical analyses*

19 Stata 14.1 was used for data management and analysis. Logistic regression was used to  
20 model the odds of being deemed appointable to anaesthetics specialty training. 'Random  
21 effects' (multi-level) models were used to control for multiple applications, which allowed  
22 for nesting of application events within doctors.  
23  
24

25 We also modelled the odds ratios of obtaining a more versus less satisfactory ARCP rating  
26 using random effects ordinal logistic regression models. This accounted for dependency of  
27 observations within individual doctors. Analyses were conducted both with and without  
28 ARCP outcomes associated with postgraduate exam failure. Trainees who fail the fellowship  
29 exam would be highly likely to receive a 'sub-optimal' ARCP rating (for example 'extended  
30 training time required'). Therefore, there was a risk that postgraduate exam achievement  
31 could be conflated with, more general, ARCP performance. Therefore we conducted  
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3 additional analyses which excluded the ARCP outcomes that were reported to have occurred  
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5 in conjunction with postgraduate exam failure. This allowed us, to some extent, to  
6  
7 disaggregate the effect of the predictors on ARCP outcomes in general from postgraduate  
8  
9 exam performance.  
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11  
12  
13 For both sets of analyses, univariable and multivariable models (using a backward stepwise  
14  
15 approach) were built. Only interaction terms that were statistically significant (at the  $p < 0.05$   
16  
17 level) were included in the final multivariable model. Missing data were managed using  
18  
19 listwise deletion. Extensive missing data (greater than 5%) were only observed for the  
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21 shortlisting and interview scores at specialty selection. Therefore, in order to evaluate the  
22  
23 potential impact of the missing data on the results a series of analyses using imputed values  
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25 for these variables were conducted as a sensitivity analysis (see the technical appendix for  
26  
27 further details).  
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## 31 32 **Results**

### 33 34 35 *Descriptive statistics*

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38 For most variables, there were relatively few missing data, except for standardised shortlist  
39  
40 and standardised interview score. 608/2782 (22%) of standardised interview scores were  
41  
42 missing, and 1,075 /2782 (39%) of standardised shortlist scores were missing. There were a  
43  
44 total of 2,782 doctors who had both recruitment and ARCP outcomes available. 52.30%  
45  
46 (1,445/2782) were males, 23.79% (661/2779) described themselves as BME ethnicity, and  
47  
48 26.21% (713/2720) were on the ACCS Anaesthesia training scheme compared to CAT.  
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51  
52 Further descriptive statistics can be seen in Table 1.  
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59 INSERT TABLE 1 HERE  
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### *Modelling recruitment outcomes*

Figure 4 displays the univariable odds ratios and 95% confidence intervals. The full results can be seen in Table S1 of the technical appendix. We can see the strongest predictor of recruitment is interview score (OR 4.70, 4.18 to 5.30,  $p < .001$ ). Shortlisting score is a positive, though weaker, predictor (OR 1.29, 1.17 to 1.43).

In terms of demographic predictors, older and/or BME trainees had significantly lower odds of being deemed appointable, whilst experience and gender were not significantly predictive (see Figure 2, and Table S1 of technical appendix). Nationals of wealthier countries were more likely to be deemed appointable (OR 1.58, 1.44 to 1.73). That is, for every \$10,000 USD GDP, a candidate's odds of success increased by approximately 60%. Similarly those who qualified in more affluent countries had higher odds of success (OR 1.86, 1.68 to 2.07).

INSERT FIGURE 4 HERE

The results from multivariable regression analysis are shown in Table 2. Since appointability of candidates is determined by their interview score, we excluded this variable from the analysis. Age at application (OR 0.94, 0.90 to 0.98), experience at application (OR 1.35, 1.25 to 1.47), BME applicants (OR 0.66, 0.49 to 0.90), and the GDP of place of qualification (OR 1.48, 1.05 to 2.10) were significant independent predictors of being deemed appointable. Experience at application was non-significant at univariable analysis, but significant at multivariable analysis. Shortlisting score was not an independent predictor of application success. No interaction terms were statistically significant.

INSERT TABLE 2 HERE

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3 Some variation in the results from imputed and non-imputed datasets were observed.

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5 Notably, when the missing selection measures were imputed, shortlisting score became a  
6  
7 significant predictor in the multivariable model, whilst experience at application became  
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9 non-significant. The full results from analyses of the imputed datasets can be found in the  
10  
11 technical appendix.  
12

### 13 14 15 *Modelling ARCP outcomes*

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18 Figure 5 displays the results from univariable analyses predicting ARCP outcomes, with and  
19  
20 without postgraduate exam failures. Full results can be found in Table S2 in the technical  
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22 appendix.  
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26 The strongest predictor of ARCP performance, when exam failures were included, was  
27  
28 training within an ACCS programme. However, this becomes non-statistically significant  
29  
30 when the effects of exam failures are eliminated. BME trainees had reduced odds of having  
31  
32 a satisfactory ARCP outcome compared to white trainees (OR 0.61, 0.52 to 0.72). Older  
33  
34 candidates (OR 0.90, 0.88 to 0.92), and those with more UK clinical experience (OR 0.79,  
35  
36 0.75 to 0.83) also, on average, had significantly less satisfactory ARCP outcomes. Once ARCP  
37  
38 outcomes associated with postgraduate exam failure were excluded from analyses, the  
39  
40 effect size reduced for all predictor variables. However, all the predictors, except for the  
41  
42 training stream group (ACCS vs CAT) remained statistically significant at the  $p < 0.05$  level.  
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48 INSERT TABLE 3 HERE  
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51 The results from the multivariable regression analyses are shown in Table 3. Shortlisting (OR  
52  
53 1.42, 1.24 to 1.62) and interview scores (OR 1.37, 1.19 to 1.57) remained independent  
54  
55 predictors of a more satisfactory ARCP outcome. For shortlisting score, for every standard  
56  
57 deviation above the mean, trainees had approximately, on average, 40% higher odds of  
58  
59 receiving a more satisfactory outcome when controlling for the influence of other variables.  
60

Age (OR 0.94, 0.91 to 0.97), UK clinical experience (OR 0.90, 0.83 to 0.97) and BME status (OR 0.69, 0.55 to 0.86) remained independently predictive of less satisfactory ARCP outcomes when exam failures were included in the analysis.

When ARCP outcomes associated with exam failure were excluded, age, experience and training stream became non-significant predictors. Interview and shortlisting scores became slightly stronger predictors after excluding ARCP outcomes associated with postgraduate exam failure. BME trainees remained significantly less likely to obtain a satisfactory ARCP outcome (OR 0.72, 0.55 to 0.93).

Results for the imputed data were similar to those from the non-imputed (original) data. The full results are contained in the technical appendix. However, in summary, using imputed data, the influence of interview scores on 'appointability' were somewhat diminished (though still statistically significant at the  $p < 0.05$  level) whilst that of the shortlisting scores increased, though did not reach statistical significance ( $p = 0.11$ ). The results for the prediction of ARCP varied relatively little between those for the imputed and non-imputed datasets.

## Discussion

This is the first national data linkage study to investigate the predictive validity of a national recruitment process in anaesthesia with measures of in-training clinical performance in training through ARCP outcomes. We found that although shortlisting scores were not predictive of appointability, they did predict ARCP success. Interview scores were not included in recruitment analysis, however they were also predictive of ARCP. At multivariate level, shortlisting and interview scores had very similar ability to predict educational success.

We did not detect a significant difference in ARCP outcomes between the two anaesthetic training streams (ACCS and CAT) once the influence of postgraduate exam failures was

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2  
3 controlled for. This is likely to reflect the differing structure of training, where CAT trainees  
4  
5 are more likely to take exams from the first year.  
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9 We observed that most demographic factors were potential confounding factors – only  
10  
11 gender and years of NHS experience were not univariable predictors of recruitment. When  
12  
13 confounders were accounted for using multivariable analyses, age at application, experience  
14  
15 at application, ethnicity, and GDP of place of qualification remained significant independent  
16  
17 predictors of appointability. Interestingly, greater experience was not a significant  
18  
19 univariable predictor of appointability but emerged as an independent predictor in the  
20  
21 multivariable model. This apparent paradox could be explained by greater years of  
22  
23 experience also possibly occurring in trainees who had encountered previous difficulty in  
24  
25 obtaining places on anaesthetics or other postgraduate programmes. It may also be a  
26  
27 mechanism of the missing data present in the study (see strengths and limitations).  
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31  
32 Poorer ARCP outcomes could be predicted by age, experience and ethnicity in both  
33  
34 univariable and multivariable models. However, age and experience were not predictive  
35  
36 once postgraduate exam outcomes were removed. Previous research has shown that junior  
37  
38 doctors aged over 29 years old are more likely to have less than satisfactory ARCP outcomes  
39  
40 across a range of specialties<sup>15</sup>. However, more recent research observed no significant  
41  
42 difference between age and satisfactory ARCP outcomes within general surgery<sup>27</sup>.  
43  
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45  
46 Postgraduate exam performance appeared more sensitive to increasing age than ARCP  
47  
48 outcomes. This could be due to older trainees being more likely to have competing family  
49  
50 responsibilities, making exam preparation more challenging, as has been suggested in  
51  
52 previous literature relating to the first year of core training<sup>15</sup>.  
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54

55  
56 GDP, of both nationality and place of qualification, were predictors of appointability,  
57  
58 although only the latter variable was an independent predictor. Both GDP variables were  
59  
60 also significant univariable predictors of ARCP outcome, but neither were independent

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3 predictors. This is reassuring in an NHS that continues to rely on international doctors to  
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5 sustain it.  
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9 Candidates, self-reporting as BME, had approximately half the odds of being deemed  
10  
11 appointable compared to those reporting White ethnicity. This effect appeared independent  
12  
13 of the other factors in the multivariable model. Similarly, BME trainees had higher odds of  
14  
15 receiving less satisfactory ARCP outcomes, even when the influence of other background  
16  
17 variables was controlled for. This effect diminished modestly, though persisted, when ARCP  
18  
19 outcomes related to postgraduate exam failures were excluded. Previous research shows  
20  
21 consistently lower performance of BME trainees in UK postgraduate medical examinations <sup>19</sup>  
22  
23 <sup>21 23</sup> and ARCP <sup>25</sup> across a number of specialties including the anaesthetic Primary FRCA <sup>28</sup>.  
24  
25 The reasons underlying such differential performance are likely to be complex, subtle and  
26  
27 multifactorial and have been much debated <sup>24 28 29</sup>.  
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33 Our findings from this national study are in line with those reported by a smaller, regional  
34  
35 one that observed that performance during recruitment correlated subsequent work place  
36  
37 based assessment ratings <sup>12</sup>.  
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#### 40 *Strengths and limitations*

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42  
43 Our study used a large national dataset related to the UK anaesthesia training selection  
44  
45 process. Moreover, we could link selection data to information relating to subsequent  
46  
47 performance and training. A further strength of this study is that many studies investigating  
48  
49 the validity of recruitment processes suffer from 'attenuating effects'. That is that low  
50  
51 performing candidates at selection do not generate an assessment score because they do  
52  
53 not get appointed. This leads to an underestimation of the recruitment process as it fails to  
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55 detect ability to filter out poorer performing candidates. However, in our study the  
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57 shortlisting and interview scores were standardized according to the performance of all  
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3 applicants, not just those accepted on to training schemes. Thus, this approach should  
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5 correct for any attenuating effects during the statistical modelling.  
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9 The data were relatively complete with the exception of the shortlisting and interview  
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11 scores. It is likely that the missing values were not 'missing completely at random'. That is,  
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13 the missing values were likely related to the values of variables that could be observed, as an  
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15 applicant could not receive an interview without a shortlist score. However, we addressed  
16  
17 this via multiple imputation as a form of sensitivity analysis. Indeed, our imputed and non-  
18  
19 imputed results were relatively similar suggesting that most missing data were randomly  
20  
21 missing, with the exception of shortlisting score. This is unsurprising given that the  
22  
23 shortlisting score determines the presence of an interview. Thus, caution must be exercised  
24  
25 when interpreting results relating to the former predictor.  
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28  
29 Another limitation of the study is that 'clinical experience' could only be measured by date  
30  
31 of registration with the GMC. Therefore years of practice outside of the UK could not be  
32  
33 accounted for. Furthermore, due to a limited numbers of overseas trainees, we were unable  
34  
35 to divide graduates by place of qualification (such as EEA graduates). We instead used the  
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37 GDP of the country of nationality or qualification and acknowledge that this has less  
38  
39 practical use than dividing by whether or not trainees were likely to have taken the  
40  
41 Professional and Linguistic Assessments Board exam. This study also focused solely on  
42  
43 anaesthetic core trainees, and the findings may not generalise to the context of higher  
44  
45 specialist training limited. Nevertheless, it is likely that those variables we observed to  
46  
47 predict success at core-training are likely to also apply to performance in subsequent higher  
48  
49 specialist training. This may be especially true of educational (as opposed to clinical)  
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51 outcomes and a high degree of continuity in academic performance in medicine is well  
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53 recognised<sup>30</sup>.  
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3 In the present study we used ARCP as one of the main outcomes. This metric lacks  
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5 granularity and is more likely to discriminate between poorer rather than better  
6  
7 performance. Data pertaining to Primary FRCA exam results was not available at the time of  
8  
9 the study. Consequently a future study using the Primary FRCA as an outcome variable is  
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11 likely to add more detail to the picture we have sketched here.  
12  
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#### 14 15 *Implications for policy and directions for future research*

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18 Our finding that both interview and shortlisting scores made similar sized, independent  
19  
20 contributions to the predict ARCP suggest that recruitment centres are correct in using both  
21  
22 scores when ranking applicants for training posts. There is an extensive literature on the  
23  
24 potential for bias of face-to-face interviews, though structuring the process, as at specialty  
25  
26 recruitment, may reduce this risk to some extent.<sup>4 12</sup> The ideal weighting for each station  
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28 should be informed by studies which are able to link data with clinical performance and our  
29  
30 results suggest that the portfolio station should make up a significant portion of the overall  
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32 selection score.  
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38 In line with previous research it was clear that certain demographic groups were less likely  
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40 to be both appointed at application to core specialty training and also more likely to receive  
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42 less satisfactory outcomes at subsequent ARCP. Certainly, for groups at risk of less  
43  
44 satisfactory ARCP outcomes it may be that additional, targeted support would be beneficial.  
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46 Such support could also be practical in nature, such as increased access to flexible working  
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48 for those with caring commitments. The most effective approaches to supporting such  
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50 individuals could be the subject of future, possibly qualitative, research.  
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54  
55 Further research could focus on follow-up of anaesthesia trainees into higher specialist  
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57 training. Performance in higher specialty trainees (STs) could be observed using subsequent  
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59 ARCP outcomes, as well as performance in the 'Final FRCA'. This exam, as with the Primary  
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3 FRCA, consists of two sections, and must be passed by the end of specialty training year 4  
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5 (ST4), in order to allow entry to more advanced specialist training (ST5 to ST7).  
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### 8 9 *Conclusions*

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12 Our findings support the effectiveness of the UK national selection process for entry to core  
13 anaesthesia training, with both shortlisting and interview scores being strongly predictive of  
14 clinical performance measured through ARCP outcomes. Demographic variables effecting  
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appointability were also associated with ARCP outcomes.

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analysis. LWP led on data analysis, produced the data visualisations, and contributed to data  
interpretation and drafting and appraising the manuscript. TG contributed to study design,  
editing and appraising the content of the manuscript. PAT led on study design and  
supervision of the project and contributed to writing, editing and appraising the content of  
the manuscript.

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1  
2  
3 **Competing interests** PAT has previously received research funding from the ESRC, the  
4  
5 EPSRC, the Department of Health for England, the UKCAT Board, and the GMC. TG has been  
6  
7 a member of the RCOA Recruitment Committee since 2010 and Chair of the committee since  
8  
9 2015. He has previously been awarded funding from the Department of Health to develop  
10  
11 and evaluate selection methods for anaesthesia training.  
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15 **Patient consent** Not required  
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19 **Ethics approval** The study relied on the analysis of de-identified routinely collected data  
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21 analysed within a 'safe haven' environment. This was confirmed in writing by the chair of the  
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23 University Of York Department of Health Sciences Ethics Committee.  
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27 **Data sharing statement** The data and associated STATA syntax used to manage and analyse  
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29 the data may be made available from the GMC on request within a safehaven environment  
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31 on an individual basis should a sufficient justification be provided.  
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#### 34 35 36 37 38 39 40 **Figure legends**

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44 Figure 1: Summary of the recruitment process into Anaesthetics Core Training.  
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47 Figure 2: Assessment pathway from medical school to the end of anaesthetics training. The  
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49 focus of our study is the Anaesthetics Core Training stage (in blue).  
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53 Figure 3: Data flowchart for the study.  
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56 Figure 4: Results from univariable logistic regression analyses for an individual being deemed  
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58 'appointable' according to each predictor variable.  
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Figure 5: Univariable ARCP results for each variable showing how including and excluding exam failures affect ARCP outcomes.

Table 1: Descriptive statistics

Variable	Mean	Standard deviation
Age on application	29.32	3.46
Experience in years on application	2.78	1.72
Standardised shortlist score	0.18	0.83
Standardised interview score	0.17	0.81
Number of jobs applied for during study period	2.25	1.31
Number of jobs deemed appointable during study period	1.49	0.93
Age at ARCP	30.13	3.33
Experience in years at ARCP	3.58	1.44
Number of ARCPs per doctor	1.83	0.93

Table 2: Results from a multivariable logistic regression model predicting the odds of a candidate being deemed appointable to core Anaesthetics training.

Variable	Odds ratio (95% CI)	p
Age at application	0.94 (0.90 to 0.98)	.008
Experience at application	1.35 (1.25 to 1.47)	<.001
BME	0.67 (0.49 to 0.90)	.008
GDP of country of PMQ	1.48 (1.05 to 2.10)	.030

Table 3: Results from multivariable ordinal logistic regression analyses predicting the odds of a satisfactory vs unsatisfactory ARCP outcome amongst anaesthetic trainees. Results excluding ARCP outcomes associated with postgraduate exam failure are in the right column.

Variable	ARCP results			
	Including exam failure		Excluding exam failure	
	Odds ratio (95% CI)	p	Odds ratio (95% CI)	p
Age at ARCP	0.94 (0.91 to 0.97)	<.001	0.97 (0.94 to 1.01)	.166
Experience at ARCP	0.90 (0.83 to 0.97)	.006	0.97 (0.88 to 1.06)	.489
Standardised interview score	1.37 (1.19 to 1.57)	<.001	1.35 (1.15 to 1.57)	<.001
Standardised shortlist score	1.42 (1.24 to 1.62)	<.001	1.28 (1.10 to 1.48)	.001
ACCS vs CT	1.43 (1.13 to 1.82)	.003	1.01 (0.78 to 1.32)	.934
BME	0.69 (0.55 to 0.86)	.001	0.72 (0.55 to 0.93)	.011

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**UK Graduates and doctors from the European Economic Area**

**Doctors with qualifications from outside of the European Economic Area**

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**Application via national ORIEL system**  
- Candidates rank geographical area (deanery), and training stream (ACCS or CAT) in which they wish to train  
- Candidates score themselves based on nationally standardised self assessment criteria which generates a shortlist score

**Selection for interview**  
Units of application utilise the shortlist score and invite applicants for interview

**Selection centre interview**  
Three stations: clinical interview, presentation and portfolio  
Two assessors independently score each station  
Self-assessment score is verified with evidence presented to assessors in the portfolio station

**Global selection score generated**  
Comprising of individual scores for each station and the shortlist self-score

**Deemed appointable or not**  
Candidates are ranked by their score and candidates over a threshold are deemed appointable or not

- Must pass the Professional and Linguistic Assessments Board exam to be eligible to apply

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**Medical School - 4 to 6 years**  
 - Medical schools set their own exams, which consist of multiple choice questions, and clinical exams



**Foundation years - 2 years**  
 - Trainees must pass an Annual Review of Competence Progression consisting of a portfolio of workplace based assessments  
 - Decision made by a panel of assessors



**Anaesthetic Core Training - 2 to 3 years**  
 - Core Anaesthetic Training - 2 years  
 - Acute Care Common Stem - 3 years



**Core Anaesthetic Training (CAT): 2 years of anaesthetic training**  
 Initial assessment of competency - within 3 months - series of workplace based assessments to prove competent to give basic anaesthetic  
 Annual Review of Competence Progression - consisting of a portfolio of workplace based assessments and assessed by a panel  
 Primary Fellowship of Royal College of Anaesthetists - usually taken within first year or beginning of second

**Acute Care Common Stem (ACCS) Anaesthetics: 1 year of acute medicine and emergency medicine in addition to 2 years of anaesthetics training**  
 - Initial assessment of competency - within first 3 months of second year - series of workplace based assessments to prove competent to give basic anaesthetic  
 - Annual Review of Competence Progression - consisting of a portfolio of workplace based assessments and assessed by a panel  
 - Primary Fellowship of Royal College of Anaesthetists - usually taken within second year or beginning of third year



**Intermediate training - 2 years**  
 - Annual Review of Competence Progression  
 - Final Fellowship of Royal College of Anaesthetics exam consisting of a written paper and clinical exam



**Higher training - 2 years**  
 - Annual Review of Competence Progression



**Advanced training - 1 years**  
 - Annual Review of Competence Progression  
 - Preparation for consultant post

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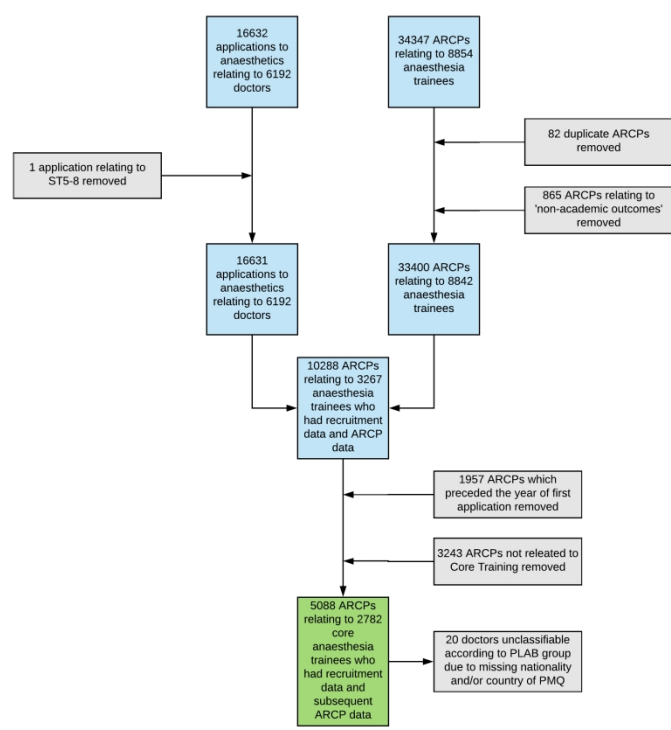


Figure 3: Data flowchart for the study.

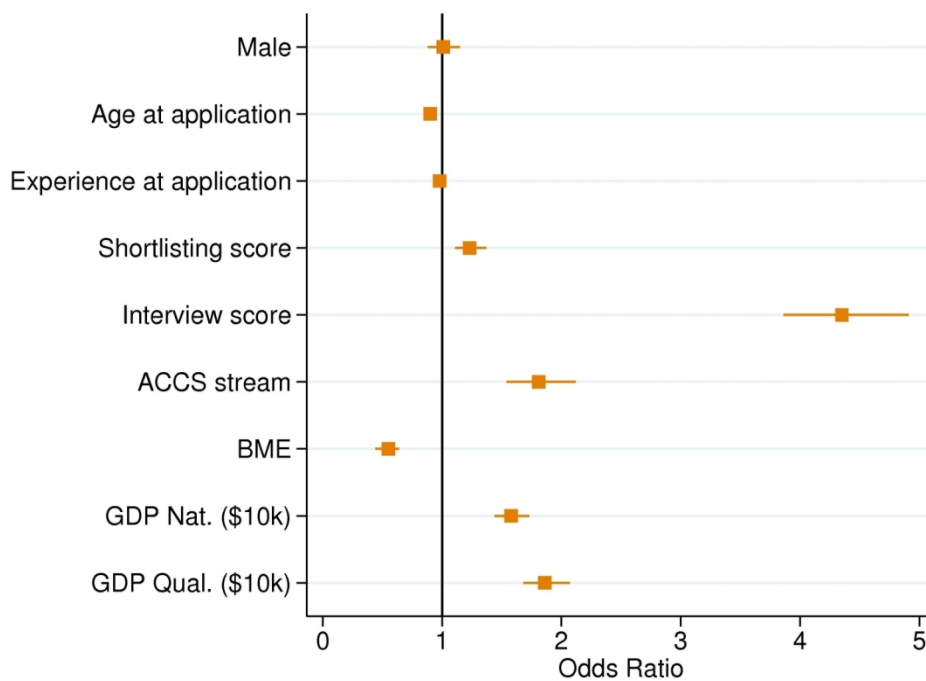


Figure 4: Results from univariable logistic regression analyses for an individual being deemed 'appointable' according to each predictor variable.

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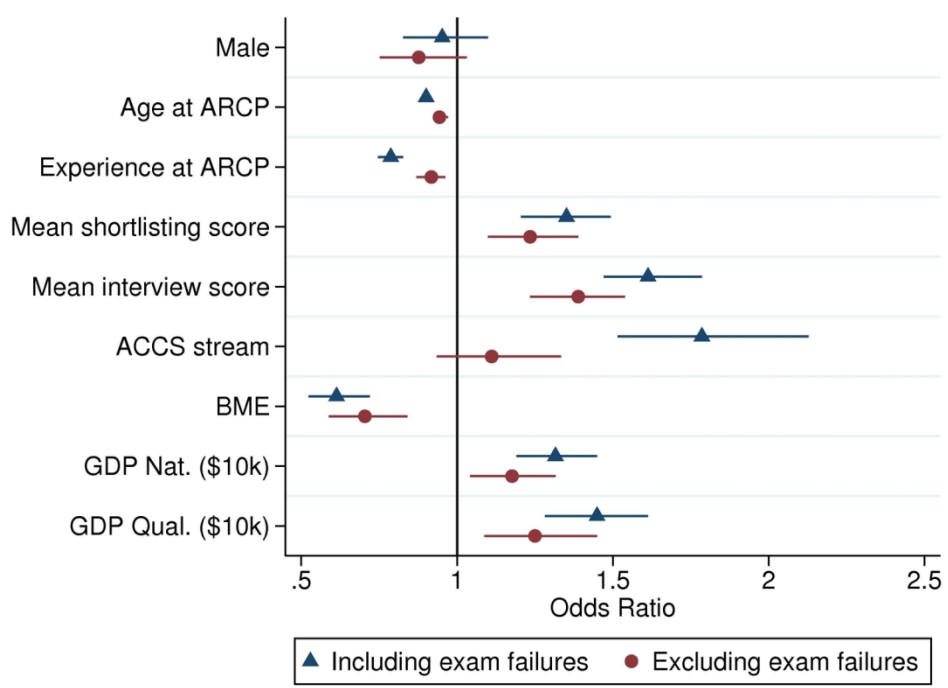


Figure 5: Univariable ARCP results for each variable showing how including and excluding exam failures affect ARCP outcomes.

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3 **Technical Appendix**  
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6 **Evaluating the recruitment process into UK anaesthesia core training: doctors'**  
7 **performance at selection and subsequent postgraduate training– a national data linkage**  
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9 **study: technical appendix**  
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15 Aslet, M., Paton, L.W., Gale, T., & Tiffin, P.A.  
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18 **A: Univariable results**  
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21 Table S1 and Table S2 detail the univariable results for being 'deemed appointable' and ARCP  
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23 outcome. These results are depicted in Figures 2 and 3 in the main paper.  
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Deemed appointable		
Variable	Odds ratio (95% CI)	p
Male gender	1.01 (0.88 to 1.15)	.92
Age at application	0.96 (0.94 to 0.97)	<.001
Experience at application	0.98 (0.95 to 1.02)	.35
Standardised shortlist score	1.23 (1.11 to 1.37)	<.001
Standardised interview score	4.35 (3.86 to 4.91)	<.001
ACCS vs CAT	1.81 (1.54 to 2.12)	<.001
BME	0.55 (0.48 to 0.64)	<.001
GDP of country of nationality (\$10k per person)	1.58 (1.44 to 1.73)	<.001
GDP of country of qualification (\$10k per person)	1.86 (1.68 to 2.07)	<.001

Table S1: Results from univariable logistic regression predicting the odds of being deemed appointable to anaesthetics training at core level.

Variable	ARCP results		Excluding exam failure	
	Including exam failure		Odds ratio (95% CI)	p
Male gender	0.95 (0.82 to 1.09)	.47	0.88 (0.75 to 1.03)	.11
Age at ARCP	0.91 (0.88 to 0.92)	<.001	0.94 (0.92 to 0.97)	<.001
Experience at ARCP	0.79 (0.75 to 0.83)	<.001	0.91 (0.87 to 0.96)	<.001
Standardised shortlist score	1.35 (1.21 to 1.50)	<.001	1.23 (1.10 to 1.39)	<.001
Standardised interview score	1.61 (1.46 to 1.78)	<.001	1.38 (1.23 to 1.54)	<.001
ACCS vs CAT	1.79 (1.52 to 2.12)	<.001	1.11 (0.93 to 1.33)	.23
BME	0.61 (0.52 to 0.72)	<.001	0.70 (0.59 to 0.83)	<.001
GDP of country of nationality (\$10k per person)	1.32 (1.19 to 1.46)	<.001	1.18 (1.05 to 1.32)	.01
GDP of country of qualification (\$10k per person)	1.44 (1.28 to 1.62)	<.001	1.25 (1.09 to 1.45)	.002

Table S2: Results from univariable analyses for a more satisfactory ARCP outcome. Analyses including exam failure are shown in the left column, and analyses excluding ARCP outcomes linked to postgraduate exam failure are shown in the right column.



## **B: Multiple imputation**

### ***Method***

As detailed in the main paper, missing data were relatively uncommon other than for shortlisting score (39% missing) and interview score (22% missing). As such, in addition to the univariable analyses and backwards stepwise multivariable regression using listwise deletion detailed in the main body of the paper, we also performed stepwise multivariable regression on multiply imputed data. We used chained equations, creating 20 imputed data sets.

This portion of the analysis can be thought of as a form of 'sensitivity analysis' for shortlisting score and interview score. That is, if the results between the imputed and non-imputed datasets vary, then this would be evidence that the absent values are 'missing not at random' (MNAR) (i.e. the missing values are neither associated with the observed data nor due to chance). If there is evidence that the absent values are MNAR then results in relation to the affected variables must be interpreted more cautiously. All analyses were performed in Stata version 14.1.

Note that in this portion of the analyses, ARCP outcome was dichotomised, with satisfactory outcomes being coded as 1, and less than satisfactory coded as 0, rather than the four-point ordinal scale used in the main text. This is due to a technical point: Stata does not support ordinal logistic regression for imputed data sets. Thus, we used the binomial outcome as discussed above. This means that the non-imputed results presented in this supplementary appendix will differ slightly from the non-imputed results presented in the main text.

Furthermore, as can be seen in Table S3 shortlisting score was not an independent predictor of being 'deemed appointable' (i.e. it was not in the final multivariable model). However, for the purposes of a sensitivity analysis, we forced inclusion in the final model of both selection variables.

## Results

### Modelling of recruitment outcomes

Variable	Deemed appointable			
	Original data		Imputed data	
	Odds ratio (95% CI)	p	Odds ratio (95% CI)	p
Age at application	0.94 (0.90 to 0.98)	.01	0.96 (0.94 to 0.98)	.001
Experience at application	1.35 (1.25 to 1.47)	<.001	1.03 (0.99 to 1.07)	.11
BME ethnicity	0.67 (0.49 to 0.90)	.01	0.66 (0.57 to 0.77)	<.001
GDP of country of qualification (\$10k per person)	1.43 (1.01 to 2.04)	.05	1.62 (1.44 to 1.82)	<.001
Standardised shortlisting score	1.16 (0.98 to 1.38)	.09	1.11 (1.01 to 1.21)	.02

Table S3: A comparison of the multivariable results for recruitment using non-imputed and imputed data. The left column shows results from the original data.

Table S3 shows the multivariable models for being deemed appointable on both non-imputed and imputed data sets. As can be seen, standardised shortlisting score became a statistically significant predictor (OR 1.11, 1.01 to 1.21) in the imputed dataset. That is, a candidate had 1.11 the odds of being appointed for each standard deviation above the mean a candidate scored. Additionally, in the imputed data, experience at application becomes non-significant.

*Modelling ARCP outcomes*

The results from the imputed data show slightly reduced effect sizes than the non-imputed results (Table S4). The most notable difference between the two sets of results is a reduction in the odds ratio for mean shortlisting score when using imputed data (Imputed result: OR 1.14, 1.03 to 1.26). A similar trend is seen when excluding ARCP outcomes related to postgraduate exam failure (Table S5). Indeed, mean shortlisting score becomes non-significant when using imputed data.

ARCP results including exam failure				
	Original data		Imputed data	
Variable	Odds ratio (95% CI)	p	Odds ratio (95% CI)	p
Age at ARCP	0.95 (0.92 to 0.98)	.001	0.94 (0.92 to 0.97)	<.001
Experience	0.91 (0.85 to 0.98)	.016	0.88 (0.84 to 0.93)	<.001
Standardised interview score	1.35 (1.18 to 1.54)	<.001	1.35 (1.22 to 1.48)	<.001
Standardised shortlist score	1.33 (1.18 to 1.51)	<.001	1.14 (1.03 to 1.26)	.015
ACCS vs CT	1.32 (1.05 to 1.66)	.019	1.26 (1.07 to 1.48)	.006
BME	0.71 (0.57 to 0.87)	.001	0.71 (0.61 to 0.83)	<.001

Table S4: A comparison of the multivariable results for ARCP outcome using non-imputed and imputed data, including ARCP outcomes related to postgraduate exam failure. The left column shows results from the original data.

ARCP results excluding exam failure				
Variable	Original data		Imputed results	
	Odds ratio (95% CI)	p	Odds ratio (95% CI)	p
Age at ARCP	0.98 (0.94 to 1.01)	.29	0.96 (0.94 to 0.99)	.002
Experience	0.97 (0.88 to 1.06)	.47	0.96 (0.91 to 1.02)	.24
Standardised interview score	1.33 (1.14 to 1.54)	<.001	1.25 (1.12 to 1.39)	<.001
Standardised shortlist score	1.23 (1.06 to 1.41)	.01	1.09 (0.96 to 1.22)	.18
ACCS v CT	0.97 (0.75 to 1.25)	.80	0.92 (0.77 to 1.09)	.34
BME	0.73 (0.57 to 0.94)	.01	0.73 (0.62 to 0.87)	<.001

Table S5: A comparison of the multivariable results for ARCP outcome using non-imputed and imputed data, excluding ARCP outcomes related to postgraduate exam failure. The left column shows results from the original data.

### **Discussion**

The imputed analyses generally show similar but slightly weaker trends for all variables for both recruitment and ARCP outcomes. The exception is shortlisting score, which becomes significant when using imputed data to predict appointability. It also shows a large reduction in effect size for ARCP outcomes, and becomes non-significant when postgraduate exam failures are excluded from the analysis. Only candidates who were successfully appointed are included in the dataset. This implies they had a relatively high shortlisting score in order to be offered an interview. This would thus reduce the ability of shortlisting scores to predict less satisfactory ARCP outcomes.

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3 A similar result can be observed when considering interview scores. Only those candidates deemed  
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5 of a high enough standard to be interviewed would subsequently receive an interview score.  
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8 Therefore, this limits the ability of this variable to predict trainees who may perform at the lower  
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10 end of the scale.  
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13 The imputed analyses for shortlisting scores show a large reduction in effect size across analyses.  
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15 This implies that the data missing for this variable are not randomly missing but systematically  
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17 missing. It may be that those deaneries who return shortlist and interview scores are generally more  
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19 efficient, which may also improve the quality of the training. This may positively skew shortlisting  
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21 scores, with those receiving higher scores also more likely have their scores returned. This would  
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23 explain why, when imputing the data with the assumption of random missingness, the effect size  
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25 observed is often reduced. Due to these differences, analyses involving shortlisting score should be  
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27 interpreted with caution.  
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